

Statistical models and likelihood functions *

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List of Definitions, Assumptions, Propositions and Theorems

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1. Statistical models

1.1 Definition STATISTICAL MODEL. A statistical model is a pair $(\mathcal{Z}, \mathcal{P})$ where \mathcal{Z} is a set of possible observations and \mathcal{P} a nonempty family of probability measures which assign probabilities to subsets of \mathcal{Z} . When the probability measures in \mathcal{P} are all defined on the same σ -algebra of events $\mathcal{A}_{\mathcal{Z}}$ in \mathcal{Z} , we shall also refer to the triplet $(\mathcal{Z}, \mathcal{A}_{\mathcal{Z}}, \mathcal{P})$ as a *statistical model*.

1.2 Definition DOMINATED MODEL. A statistical model $(\mathcal{Z}, \mathcal{A}_{\mathcal{Z}}, \mathcal{P})$ is dominated if all the probability measures in \mathcal{P} have a density with respect to the same measure μ on \mathcal{Z} . μ is called the *dominating measure* and we say that $(\mathcal{Z}, \mathcal{P})$ is μ -dominated.

1.3 Definition HOMOGENEOUS MODEL. A statistical model $(\mathcal{Z}, \mathcal{A}_{\mathcal{Z}}, \mathcal{P})$ is homogeneous if it is dominated and the dominating measure μ can be chosen so that the densities are all strictly positive.

1.4 Definition PARAMETRIC MODEL. A statistical model $(\mathcal{Z}, \mathcal{P})$ is said to be parametrized by the elements of a nonempty set Θ if the set \mathcal{P} of probability

measures has the form

$$\mathcal{P} = \{P_\theta : \theta \in \Theta\} .$$

If the set Θ is a subset of \mathbb{R}^p or we can define a one-to-one transformation between Θ and the elements of a subset of \mathbb{R}^p , we say that $(\mathcal{Z}, \mathcal{P})$ is a parametric model. Otherwise, the model $(\mathcal{Z}, \mathcal{P})$ is said to be nonparametric.

1.5 Definition FUNCTIONAL PARAMETER. A functional parameter on a statistical model $(\mathcal{Z}, \mathcal{P})$ is an application

$$g : \mathcal{P} \rightarrow \Theta$$

which assigns to each element $P \in \mathcal{P}$ a parameter $\theta = g(P) \in \Theta$, where Θ is a nonempty set (the parameter space).

Functional parameters allow one to associate parameters with the distributions of parametric or non-parametric models. The mean, variance, median, etc., of a probability distribution may all be interpreted as functional parameters.

2. Identification

Let $(\mathcal{Z}, \mathcal{A}_{\mathcal{Z}}, \mathcal{P})$ a statistical model such that $\mathcal{P} = \{P_{\theta} : \theta \in \Theta\}$.

2.1 Definition IDENTIFICATION OF A PARAMETER VALUE. We say that a parameter value $\theta_1 \in \Theta$ is identifiable if there is no other value $\theta_2 \in \Theta$ such that $P_{\theta_1} = P_{\theta_2}$.

2.2 Definition IDENTIFICATION OF A MODEL. We say that the model $(\mathcal{Z}, \mathcal{A}_{\mathcal{Z}}, \mathcal{P})$ is identifiable if all the elements of Θ are identifiable.

2.3 Definition IDENTIFICATION OF A PARAMETRIC FUNCTION. Let $\psi : \theta \rightarrow \Psi$ be a function of θ . We say that the function $\psi(\theta)$ is identifiable if

$$\psi(\theta_1) \neq \psi(\theta_2) \Rightarrow P_{\theta_1} \neq P_{\theta_2}, \forall \theta_1, \theta_2 \in \Theta$$

or, equivalently,

$$P_{\theta_1} = P_{\theta_2} \Rightarrow \psi(\theta_1) = \psi(\theta_2), \forall \theta_1, \theta_2 \in \Theta.$$

2.4 Definition LOCAL IDENTIFICATION. Suppose the set Θ has a set of neighborhoods defined on it (a topology). Then we say that a parameter value $\theta_1 \in$

Θ is locally identifiable if there is a neighborhood $V(\theta_1)$ of θ_1 such that

$$\theta_2 \in V(\theta_1) \text{ and } \theta_2 \neq \theta_1 \Rightarrow P_{\theta_1} \neq P_{\theta_2} .$$

3. Likelihood and score functions

3.1 Definition LIKELIHOOD FUNCTION. Let $(\mathcal{Z}, \mathcal{P})$ be a statistical model which satisfies the following assumptions:

- (A1) $(\mathcal{Z}, \mathcal{P})$ is a μ -dominated model;
- (A2) $\mathcal{P} = \{P_\theta : \theta \in \Theta \subseteq \mathbb{R}^p\}$;
- (A3) $L(z; \theta)$, $z \in \mathcal{Z}$, is the density function (with respect to μ) associated with P_θ .

The density function $L(z; \theta)$ viewed as a function of θ is called the likelihood function of model $(\mathcal{Z}, \mathcal{P})$. The symbol $E_\theta(\cdot)$ refers to the expected value with respect to θ (provided it exists) :

$$E_\theta[h(Z)] = \int_{\mathcal{Z}} h(z) dP_\theta(z) = \int_{\mathcal{Z}} h(z) L(z; \theta) d\mu(z) .$$

The vector Z often has the form

$$Z = (Y_1', Y_2', \dots, Y_n)'$$

where $Y_t \in \mathbb{R}^m$ is an “individual” observation vector and $\theta = (\theta_1, \theta_2, \dots, \theta_p)' \in \Theta$. Usually, the density $L(z; \theta)$ is written in the form

$$L(z; \theta) = \prod_{t=1}^n f_t(z; \theta) \equiv L_n(z; \theta) \quad (3.1)$$

where $f_t(z; \theta)$ is a density for an “individual observation”. $f_t(z; \theta)$ usually has one of the following forms :

$$f_t(z; \theta) = f(y_t; \theta) , \quad y_t \in \mathbb{R}^m \quad (3.2)$$

$$f_t(z; \theta) = f(y_t | x_t; \theta) \quad (3.3)$$

where x_t is a $k \times 1$ vector of conditioning variables (“explanatory variables”) and $f(y_t; \cdot)$ is the density function of y_t (given x_t) as a function of the parameter vector θ , or

$$L_t(z; \theta) = f(y_t | \bar{y}_{t-1}, x_t; \theta) \quad (3.4)$$

where $\bar{y}_{t-1} = (\bar{y}_0, y_1, \dots, y_{t-1})'$ is a vector of past values of y and \bar{y}_0 is a vector of “initial conditions”.

3.2 Definition SCORE FUNCTION. *Under the assumption (A1) to (A3), suppose also that:*

(A4) Θ is an open set in \mathbb{R}^p ;

(A5) $\partial L(z; \theta) / \partial \theta$ exists, $\forall z \in \mathcal{Z}$, $\forall \theta \in \Theta$;

(A6) $L(z; \theta) > 0$, $\forall z \in \mathcal{Z}$, $\forall \theta \in \Theta$;

(A7)
$$\int_{\mathcal{Z}} \frac{\partial}{\partial \theta} [L(z; \theta)] d\mu(z) = \frac{\partial}{\partial \theta} \left[\int_{\mathcal{Z}} L(z; \theta) d\mu(z) \right].$$

Then the function

$$S(z; \theta) = \frac{\partial}{\partial \theta} [\ln L(z; \theta)] , \theta \in \Theta , z \in \mathcal{Z} ,$$

is called the score function associated with the likelihood $L(z; \theta)$.

3.3 Proposition MEAN OF A SCORE. *Under the assumptions (A1) to (A7), we have :*

$$E_{\theta} [S(Z; \theta)] = \int_{\mathcal{Z}} S(z; \theta) L(z; \theta) d\mu(z) = 0 .$$

3.4 Definition INFORMATION MATRIX. *In addition to (A1) to (A7), suppose also that:*

(A8) $S(Z; \theta)$ has finite second moments with respect to $P_\theta, \forall \theta \in \Theta$.

Then, the covariance matrix of $S(Z; \theta)$,

$$\begin{aligned} I(\theta) &= V_\theta [S(Z; \theta)] = E_\theta [S(Z; \theta) S(Z; \theta)'] \\ &= \int_{\mathcal{Z}} S(z; \theta) S(z; \theta)' L(z; \theta) d\mu(z) \end{aligned}$$

is called the Fisher information matrix associated with $L(z; \theta)$.

3.5 Proposition INFORMATION MATRIX IDENTITY. Under the assumptions (A1) to (A8), suppose also that:

(A9) $\frac{\partial^2 L(z; \theta)}{\partial \theta \partial \theta'}$ exists, $\forall z \in \mathcal{Z}, \forall \theta \in \Theta$;

(A10) $\forall \theta \in \Theta$,

$$\int_{\mathcal{Z}} \frac{\partial^2 L(z; \theta)}{\partial \theta_i \partial \theta_j} d\mu(z) = \frac{\partial^2}{\partial \theta_i \partial \theta_j} \left[\int_{\mathcal{Z}} L(z; \theta) d\mu(z) \right].$$

Then

$$I(\theta) = E_\theta \left[-\frac{\partial^2 \ln L(Z; \theta)}{\partial \theta \partial \theta'} \right], \forall \theta \in \Theta.$$

4. Efficiency bounds

4.1 Definition REGULAR ESTIMATOR. *Under the assumptions (A1) to (A5), an estimator $T(Z)$ of some function $\psi(\theta) \in \mathbb{R}^q$ is regular if it satisfies the following properties:*

- (a) $T(Z)$ has finite second moments;
- (b) $\int_{\mathcal{Z}} T(z) L(z; \theta) d\mu(z)$ is differentiable with respect to θ ;
- (c) $\frac{\partial}{\partial \theta} \int_{\mathcal{Z}} T(z) L(z; \theta) d\mu(z) = \int_{\mathcal{Z}} T(z) \frac{\partial}{\partial \theta} [L(z; \theta)] d\mu(z)$, for all $\theta \in \Theta$.

4.2 Theorem FRÉCHET-DARMOIS-CRAMER-RAO BOUND. *Let the assumptions (A1) to (A8) hold, let $\psi(\theta) \in \mathbb{R}^q$ be a differentiable function of θ , and suppose that*

- (A11) *the information matrix $I(\theta)$ is positive definite, $\forall \theta \in \Theta$.*

If $E_{\theta}[T(Z)] = \psi(\theta)$, $\forall \theta \in \Theta$, then the difference

$$V_{\theta}[T(Z)] - P(\theta) I(\theta)^{-1} P(\theta)'$$

is positive semi-definite for all $\theta \in \Theta$, where $P(\theta) = \partial \psi(\theta) / \partial \theta'$.

4.3 Remark If $\psi(\theta) = \theta$, this means that $V_{\theta}[T(Z)] - I(\theta)^{-1}$ is positive semi-definite.

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